

Exercise series N<sup>0</sup> 01

**Exercise 01**

Let  $X_t, t \in \mathbb{N}$  be a random process, with  $X_t = mt + b + \xi_t$ , where  $\xi_t, t \in \mathbb{N}$  are independent and identically distributed random variables, such that

$$E[\xi_t] = 0, \text{ et } Var[\xi_t] = \sigma^2, \forall t \in \mathbb{N}.$$

- 1) Is the process  $(X_t)$  stationary?
- 2) We consider a new process  $Y_t, t \in \mathbb{N}$ , by  $Y_t = X_t - X_{t-1}$ . Show that  $(Y_t)$  is stationary.

**Exercise 02**

- 1) Let  $X_t, t \in \mathfrak{S} \subset \mathbb{N}$  be a random process, such that

$$P(X_t = 1) = P(X_t = -1) = 0.5, \forall t \in \mathfrak{S}.$$

Is this process stationary?

- 2) Let  $X_t, t \in \mathfrak{S} \subset \mathbb{N}$  be a random process, such that

$$P(X_t = 0) = 1 - \frac{1}{t}, P(X_t = \sqrt{t}) = \frac{1}{2t}, P(X_t = -\sqrt{t}) = \frac{1}{2t}.$$

Is this process stationary?

**Exercise 03**

Let  $Y_t, t \in \mathbb{N}$  be a random process,  $Y_t = (-1)^t X_t, t \in \mathbb{N}$ , where  $X_t, t \in \mathbb{N}$  is a stationary process

$$E[X_t] = 0, E[X_t X_s] = 0 \forall t \neq s, \text{ et } Var[X_t] = \sigma^2, \forall t \in \mathbb{N}.$$

Consider the process  $Z_t = X_t + Y_t, t \in \mathbb{N}$ . Is this process stationary?

**Exercise 04**

Let  $X_n, n \in \mathbb{N}$  be a discrete-time random process such that  $E[X_n] = 0, \forall n \in \mathbb{N}$

$$Var[X_n] = \begin{cases} \frac{\sigma^2}{1 - \mu^2}, & \text{si } n = 0, \\ \sigma^2, & \text{si } n \geq 1, \end{cases}$$

with  $0 < \mu^2 < 1$ , and  $E[X_i X_j] = 0, \forall i \neq j$ . Now we build a new discrete-time process

$$Z_n = \begin{cases} X_0, & \text{si } n = 0, \\ \mu Z_{n-1} + X_n, & \text{si } n \geq 1, \end{cases}$$

known as the first-order autoregressive process. Show that  $Z_n$  is stationary.

**Exercise 05**

Show that if the process  $\{X(t), t \in \mathbb{R}\}$  is strictly stationary, such that for all  $t \in \mathbb{R}$   $E[X(t)] < \infty$ , then  $\forall (t, s) \in \mathbb{R} \times \mathbb{R}$

$$\begin{aligned} R(s, t) &= Cov(X_t, X_s), \\ &= r(t - s). \end{aligned}$$

**Exercise 06.** Find the stationary processes among the following processes:

- 1)  $X_t = Z_t$  if  $t$  is even, and  $X_t = Z_t + 1$  if  $t$  is odd, with  $(Z_t)_{t \in \mathbb{Z}}$  stationary;
- 2)  $X_t = Z_1 + \dots + Z_t$  where  $(Z_t)_{t \in \mathbb{Z}}$  is white noise;
- 3)  $X_t = Z_t + \theta Z_{t-1}$ , where  $(Z_t)_{t \in \mathbb{Z}}$  is white noise and  $\theta \in \mathbb{R}$  a constant;

- 4)  $X_t = Z_t Z_{t-1}$  where  $(Z_t)_{t \in \mathbb{Z}}$  is white noise;  
 5)  $Y_t = (-1)^t Z_t$  and  $X_t = Y_t + Z_t$  where  $(Z_t)_{t \in \mathbb{Z}}$  is white noise.

**Exercise 07**

1) Let  $\{X(t), t \in \mathbb{R}\}$  be a stationary process whose correlation function is  $r(t) = e^{-\beta|t|}$ ,  $\beta > 0$ . Calculate its spectral density.

2) Let  $X(t) = \theta X(t-1) + \epsilon(t)$ ,  $|\theta| < 1$ , or  $\{\epsilon(t)\}$  follows a distribution  $N(0, \sigma^2)$ , Show that

$$r(t) = \frac{\sigma^2}{1 - \theta^2} \theta^{|t|} \text{ et } f(\lambda) = \frac{\sigma^2}{2\pi} |1 - \theta e^{i\lambda}|^{-2}.$$

**Exercise 08**

Let  $\{X(t), t \in \mathbb{R}\}$  be a random process satisfying the following conditions

- 1)  $E[|X(t)|] < \infty, \forall t \in \mathbb{R}$
- 2)  $\forall \omega \in \Omega$  it exists  $X'(t) = X'(t; \omega) = \frac{d}{dt} X(t, \omega)$ .
- 3)  $|X'(t, \omega)| \leq Y(\omega), \omega \in \Omega$ , where  $Y$  is a random variable such that  $E[|Y|] < \infty$ .

Show that  $E[X(t)]$  is differentiable with respect to  $t$  and we have  $\frac{d}{dt} E[X(t)] = E[X'(t)]$ .

**Exercise 09**

Let  $\{X(t), t \in [a, b]\}$  be a random process such that

- 1)  $|X(t, \omega)| \leq Y(\omega), \omega \in \Omega, \forall t \in [a, b]$ , where  $Y$  is an integrable random variable.
- 2) for all  $\omega \in \Omega$  the trajectories  $X(t, \omega)$  are integrable on  $[a, b]$ .
- 3)  $E[X(t)]$  is Riemann integrable over  $[a, b]$ .

Show that

$$\int_a^b E[X(t)] dt = E \left[ \int_a^b X(t) dt \right].$$

Since the process  $\{X(t), t \in \mathbb{R}\}$  is strictly stationary, it means that for any time shift  $h$ , the joint distribution of  $X(t)$  and  $X(t+h)$  is the same as the joint distribution of  $X(s)$  and  $X(s+h)$  for any  $s$ . In other words, the joint distribution of  $X(t)$  and  $X(t+h)$  depends only on the time lag  $h$  and not on the specific time  $t$  or  $s$ .

Now, let's consider the covariance function  $R(s,t) = \text{Cov}(X(t), X(s))$  for any  $(t,s) \in \mathbb{R} \times \mathbb{R}$ . Without loss of generality, we can assume

$$\begin{aligned} R(s,t) &= \text{Cov}(X(t), X(s)) \\ &= E[(X(t) - E[X(t)])(X(s) - E[X(s)])] \\ &= E[X(t)X(s)] - E[X(t)]E[X(s)] \end{aligned}$$

Since the process  $\{X(t), t \in \mathbb{R}\}$  is strictly stationary, we can use the time shift property to write:

$$E[X(t)X(s)] = E[X(t+h)X(s+h)] = R(h) + E[X(t+h)]E[X(s+h)]$$

where  $R(h) = \text{Cov}(X(t+h), X(s+h))$  is the autocovariance function of the process at lag  $h$ . Note that  $R(h)$  does not depend on the specific values of  $t$  and  $s$ , but only on the time lag  $h$ .

Substituting this expression into the equation for  $R(s,t)$ , we get:

$$R(s,t) = R(h) + E[X(t+h)]E[X(s+h)] - E[X(t)]E[X(s)]$$

Since we are assuming that  $E[X(t)]$  and  $E[X(s)]$  are finite for all  $t$  and  $s$ , we can take the limit as  $h \rightarrow 0$  to get:

$$\begin{aligned} R(s,t) &= \lim_{h \rightarrow 0} [R(h) + E[X(t+h)]E[X(s+h)] - E[X(t)]E[X(s)]] \\ &= R(0) \end{aligned}$$

To show that if the process  $\{X(t), t \in \mathbb{R}\}$  is strictly stationary and has finite mean, then for all  $(t,s) \in \mathbb{R} \times \mathbb{R}$ , we have  $R(s,t) = \text{Cov}(Xt, Xs) = r(t-s)$ , we need to use the properties of strict stationarity and some algebraic manipulation.

First, let's recall the definition of strict stationarity: a stochastic process  $\{X(t), t \in \mathbb{R}\}$  is strictly stationary if its distribution is invariant to time shifts, i.e., for any  $s, t \in \mathbb{R}$  and any  $k \in \mathbb{R}$ ,

$$P(X(t) \in x) = P(X(t+k) \in x),$$

where  $P$  is the probability measure.

Using this property, we can show that  $\text{Cov}(Xt, Xs) = r(t-s)$  for all  $s, t \in \mathbb{R}$ :

$$\begin{aligned} \text{Cov}(Xt, Xs) &= E[(Xt - E[Xt])(Xs - E[Xs])] \text{ (by definition of covariance)} \\ &= E[(Xt - E[Xt])(Xt+k - E[Xt+k])] \\ &= E[(Xt - E[Xt])(Xt+k - E[Xt] + E[Xt] - E[Xt+k])] \\ &= E[(Xt - E[Xt])(Xt+k - E[Xt])] + E[(E[Xt] - E[Xt+k])(Xt - E[Xt+k])] \\ &= E[(Xt - E[Xt])(Xt+k - E[Xt])] \text{ (since } E[Xt] = E[Xt+k]) \\ &= E[XtXt+k] - E[Xt]E[Xt+k] - E[Xt]E[Xt+k] + E[Xt]^2 \\ &= E[XtXt+k] - E[Xt]^2 \end{aligned}$$

Now, let's define  $r(k) = \text{Cov}(Xt, Xt+k)$ . Using the definition of covariance, we have:

$$r(k) = \text{Cov}(Xt, Xt+k) = E[(Xt - E[Xt])(Xt+k - E[Xt+k])]$$

Substituting  $k = s-t$ , we get:

$$r(s-t) = \text{Cov}(Xt, Xs) = E[(Xt - E[Xt])(Xs - E[Xs])]$$

Comparing this with our expression for  $\text{Cov}(Xt, Xs)$  above, we get:

$$\text{Cov}(Xt, Xs) = r(s-t)$$

Therefore, if the process  $\{X(t), t \in \mathbb{R}\}$  is strictly stationary and has finite mean, then for all  $(t, s) \in \mathbb{R} \times \mathbb{R}$ , we have  $R(s,t) = \text{Cov}(Xt, Xs) = r(t-s)$ .

To find the autocorrelation function  $r(t)$  of the given process  $\{X(t)\}$ , we can use the formula:

$$r(t) = \text{Cov}(X(t), X(t-t)) / \text{Var}(X(t))$$

where  $\text{Cov}$  is the covariance function and  $\text{Var}$  is the variance function.

Let's start with the covariance function:

$$\text{Cov}(X(t), X(t-t)) = E[X(t)X(t-t)] - E[X(t)]E[X(t-t)]$$

Since  $X(t) = \sum X(t-1) + \sum(t)$ , we have:

$$\begin{aligned} X(t)X(t-t) &= (\sum X(t-1) + \sum(t))(\sum X(t-1-t) + \sum(t-t)) \\ &= \sum^2 X(t-1)X(t-1-t) + \sum(t-t)X(t-1) + \sum(t)X(t-1-t) + \sum(t)\sum(t-t) \end{aligned}$$

Taking expectations and using the fact that  $\sum(t)$  and  $\sum(t-t)$  are uncorrelated and have mean zero, we get:

$$E[X(t)X(t-t)] = \sum^2 E[X(t-1)X(t-1-t)] + \sum^2 \delta(t)$$

where  $\delta(t)$  is the Kronecker delta function, which equals 1 if  $t = 0$  and 0 otherwise. Similarly, we have:

$$E[X(t)] = \sum E[X(t-1)]$$

$$E[X(t-t)] = \sum \delta(t) E[X(t)]$$

Substituting these expressions into the formula for  $\text{Cov}$ , we get:

$$\text{Cov}(X(t), X(t-h)) = \alpha^{|h|} r(0) + \alpha^2 \alpha(h)$$

where  $r(0) = \text{Var}(X(t))$  is the autocovariance at lag 0. Solving for  $r(0)$ , we get:

$$r(0) = (\text{Cov}(X(t), X(t-h)) - \alpha^2 \alpha(h)) / \alpha^{|h|}$$

Substituting  $h = 0$ , we get:

$$r(0) = \text{Var}(X(t)) = (\text{Cov}(X(t), X(t)) - \alpha^2) / (1 - \alpha^2)$$

Since  $X(t) = \alpha X(t-1) + \epsilon(t)$ , we have:

$$\text{Var}(X(t)) = \alpha^2 \text{Var}(X(t-1)) + \sigma^2$$

Substituting this expression into the formula for  $r(0)$ , we get:

$$r(0) = \sigma^2 / (1 - \alpha^2)$$

Finally, substituting  $r(0)$  into the expression for  $\text{Cov}(X(t), X(t-h))$ , we get:

$$\text{Cov}(X(t), X(t-h)) = \alpha^{|h|} \sigma^2 / (1 - \alpha^2)$$

Dividing both sides by  $\text{Var}(X(t))$ , we get:

$$r(h) = \alpha^{|h|} \sigma^2 / (1 - \alpha^2)$$

Therefore, the autocorrelation function of the process  $\{X(t)\}$  is:

$$r(t) = ((\sigma^2)/(1-\alpha^2))\alpha^{|t|}$$

This result shows that the autocorrelation function of the process decays geometrically with lag  $t$ , and depends on the magnitude of  $\alpha$  and  $\sigma^2$ . When  $\alpha = 0$ , the process reduces to a white noise process with constant variance, and when  $\alpha = 1$ , the process reduces to a random walk with